

# Measuring Volatility

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Marcel Visser

m.p.visser@uva.nl

<http://staff.science.uva.nl/~marvisse/>

University of Amsterdam, Korteweg-de Vries Institute for Mathematics

# Goal

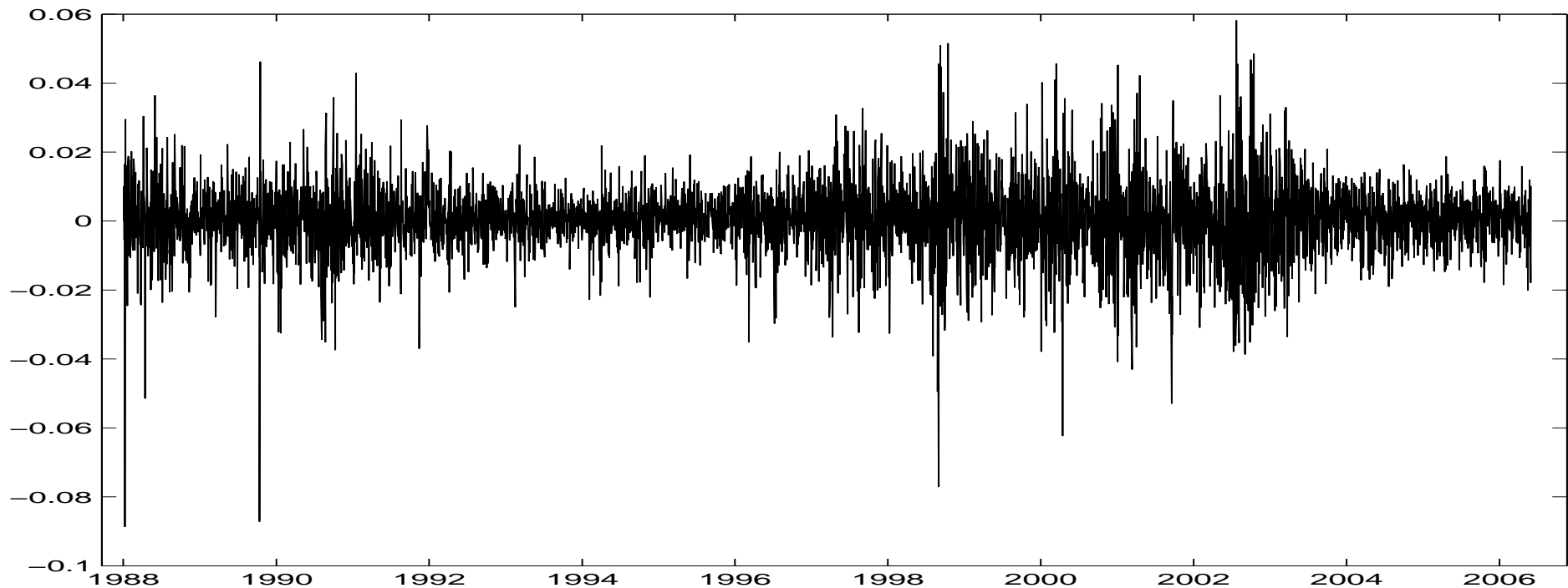
To find a good proxy for the daily scale factor  $s_n$  by using intraday data.

What does this mean?

# Introduction

DATA: 18 years of S&P 500 tick data.

Daily returns  $r_n$ : volatility clustering



# Introduction

Empirical stylized facts of daily asset returns:

- $r_n$  roughly unpredictable
- $r_n$  heavy tailed
- $|r_n|$  partly predictable (slowly decaying autocorrelations)

# Introduction

Discrete time, close-to-close return models typically have a multiplicative form

$$r_n = s_n Z_n,$$

where

- $s_n$  : the daily volatility (a scale factor  $>0$ )
- $Z_n$  : the iid innovations

A particular example for the volatility dynamics is Garch(1,1):

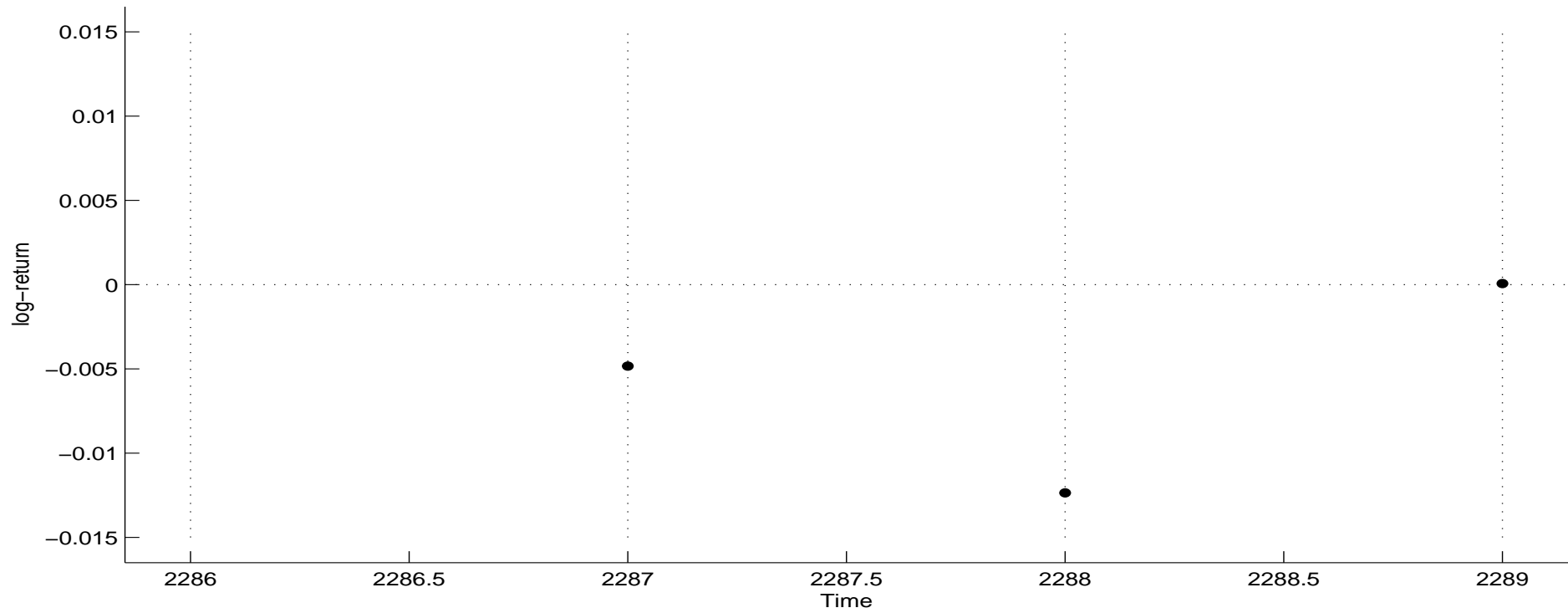
$$s_n^2 = \kappa + \alpha r_{n-1}^2 + \beta s_{n-1}^2.$$

Other models include:

stochastic volatility, long memory, Markov switching

# Introduction

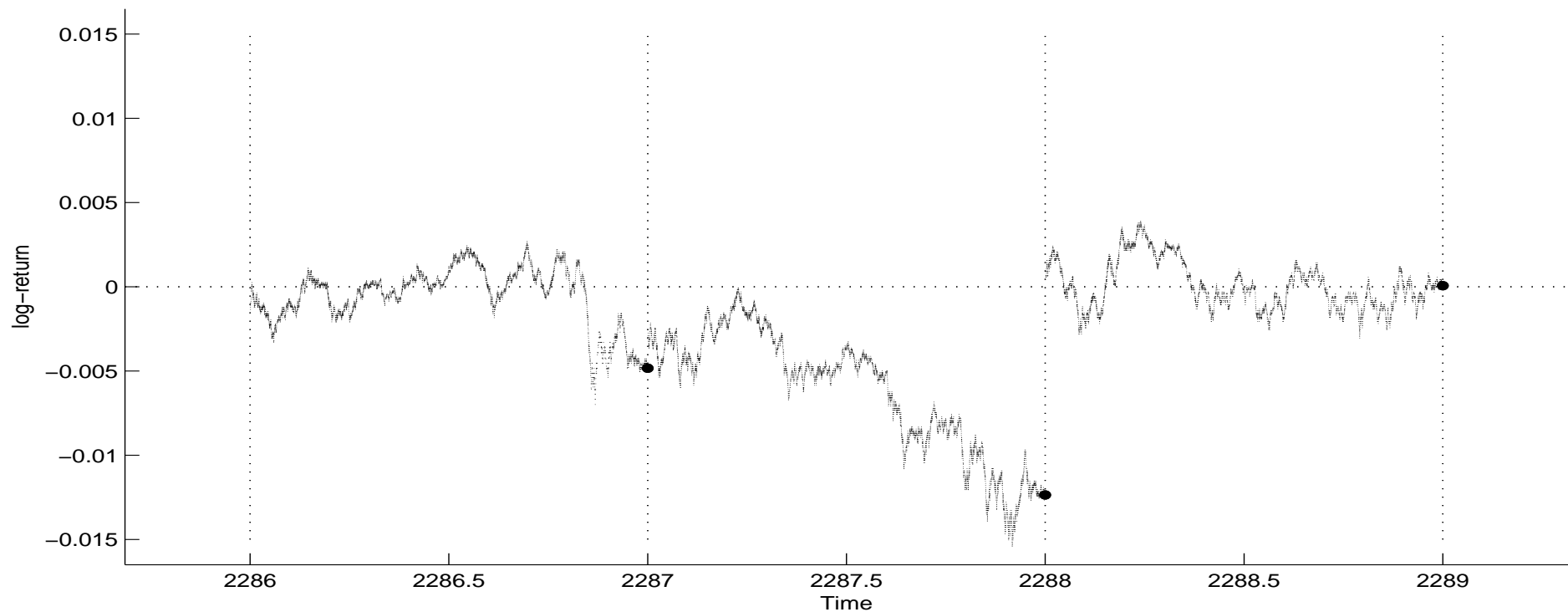
Three close-to-close returns ...



# Introduction

Three daily returns ...

and the intraday log-return processes



How to use intraday information? (these 3 days: 14 thousand price ticks)

# Introduction

We will use the intraday data to construct measurements (or proxies) for  $s_n \dots$

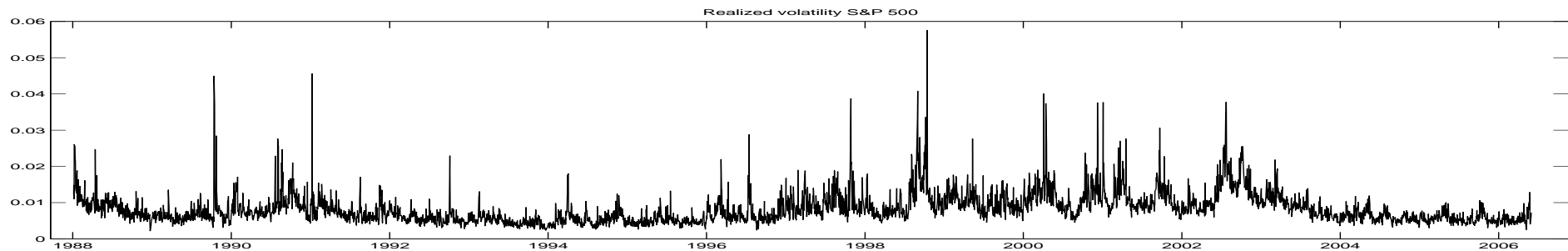
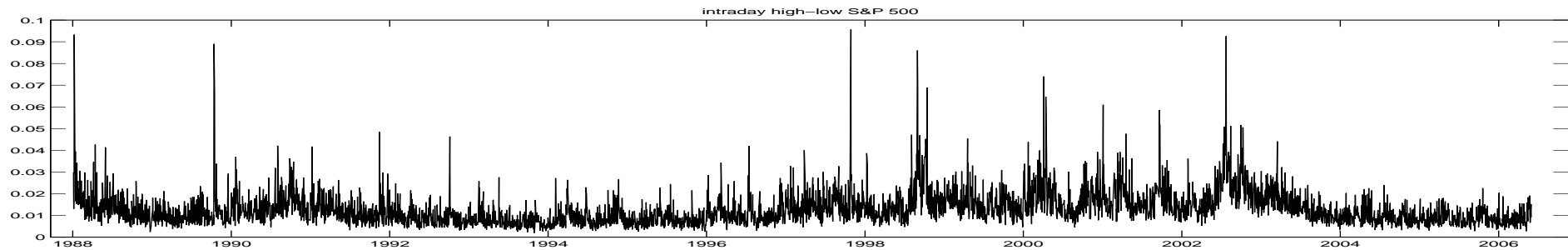
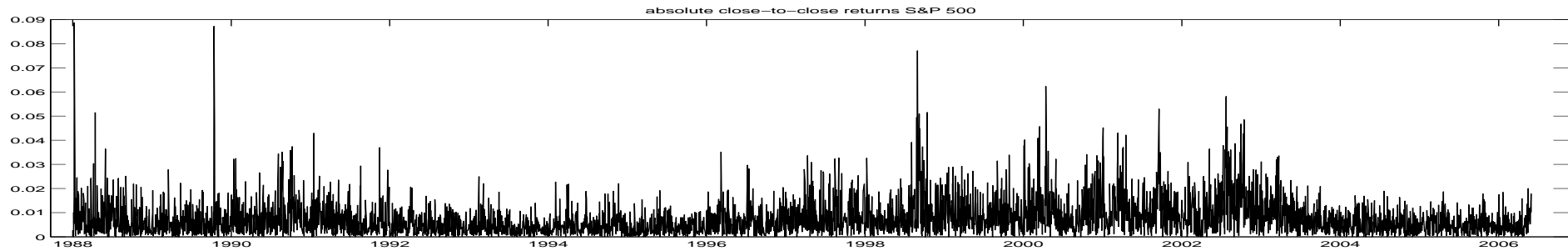
The discrete time model  $r_n = s_n Z_n$  needs proxies for

- estimation
- forecast evaluation

Which of the following three 'proxies' for  $s_n$  is best?

# Three Proxies

absolute returns vs high-low vs realized volatility



# Literature

## Measuring Volatility:

Quadratic Variation	scale factor $s_n$
<ul style="list-style-type: none"><li>• realized variance (sum of squared intraday returns; ABDL, 1998, 2003; BNS, 2002)</li><li>• realized (bi)power variation (BNS, 2004)</li><li>• realized range (MD, 2007)</li></ul>	<ul style="list-style-type: none"><li>• <math> r_n </math></li><li>• high-low (under assumption of Brownian Motion: Parkinson, 1980; ABD, 2002)</li><li>• More general: ?</li></ul>

# Step 1/3: Intraday Extension of $r_n = s_n Z_n$

Let the intraday return process  $R_n(\vartheta)$ ,  $\vartheta \in [0, 1]$  be given by

$$R_n(\vartheta) = s_n \Psi_n(\vartheta),$$

where  $\Psi$  is an arbitrary cadlag process. So,

$$r_n = R_n(1) = s_n \Psi_n(1).$$

Moreover:

- $(\Psi_n)$  iid.
- $s_n$  and  $\Psi_n$  independent, for each  $n$ .

# Step 2/3: Comparing Proxies

Definition. Let  $H$  be a positively homogeneous functional:  
 $H(R_n) \geq 0$  and

$$H(\alpha R_n) = \alpha H(R_n) \quad \alpha \geq 0$$

Then  $H(R_n)$  is a proxy for  $s_n$  and

$$H(R_n) = s_n H(\Psi_n).$$

This covers the usual proxies:

high-low, realized volatility, realized range

# Step 2/3: Comparing Proxies

Applying logarithms gives an additive measurement equation:

$$\log(H(R_n)) = \log(s_n) + U_n.$$

The errors  $U_n = \log(H(\Psi_n))$  are iid.

A good proxy should add only little variation to the variation that is in  $s_n \Rightarrow$

compare proxies by their measurement variance:

$$\lambda^2 = \text{var}(U_n).$$

# Step 2/3: Comparing Proxies

But:  $\lambda^2$  cannot be estimated! However, by independence:

$$\text{var}(\log(H^{(i)}(R_n))) = \text{var}(\log(s_n)) + (\lambda^{(i)})^2.$$

So, for two functionals  $H^{(1)}, H^{(2)}$  :

$$(\lambda^{(1)})^2 \leq (\lambda^{(2)})^2$$

$$\Leftrightarrow$$

$$\text{var}(\log(H^{(1)}(R_n))) \leq \text{var}(\log(H^{(2)}(R_n)))$$

$\Rightarrow$  Rank proxies by the variance of their logarithm.

# Step 3/3: Combining Proxies

Given  $H^{(1)}, H^{(2)}$  consider the geometric combination

$$H^{(w)}(R_n) \equiv (H^{(1)}(R_n))^{w_1} (H^{(2)}(R_n))^{w_2} \quad w_1 + w_2 = 1$$

$\Rightarrow$  Find  $w = (w_1, w_2)$  such that  $\text{var}(\log(H^{(w)}(R_n)))$  is minimal.

# Step 3/3: Combining Proxies

Define the log-proxy covariance  $\Lambda_p$  :

$$\Lambda_p \equiv \text{Cov} \begin{pmatrix} \log(H^{(1)}(R_n)) \\ \log(H^{(2)}(R_n)) \end{pmatrix}$$

Then,

$$\arg \min_w \text{var}(\log(H^{(w)}(R_n))) \quad \text{s.t.} \quad \sum w_i = 1.$$

$\Leftrightarrow$

$$\arg \min_w w' \Lambda_p w \quad \text{s.t.} \quad \sum w_i = 1.$$

# Step 3/3: Combining Proxies

Similar to Markowitz minimal variance portfolio:

$$w^* = \frac{\Lambda_p^{-1} \iota}{\iota' \Lambda_p^{-1} \iota}$$

$\Lambda_p$  can be estimated.

# Empirical Analysis

S&P 500 tick data 1988–2006 (smaller  $PV$  corresponds to better proxy)

full sample

name	PV
RV5	0.064
RV10	0.080
RV15	0.089
RV20	0.100
RV30	0.117
abs-r	0.611
hl	0.161
maxar2	0.118
RAV5	0.058
RAV10	0.072
RVHL10	0.053
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# Empirical Analysis

Same ranking for all subsamples

	full	1st	2nd	3rd	4th
name	ranking	ranking	ranking	ranking	ranking
<b>RV5</b>	4	4	4	4	4
<b>RV10</b>	6	6	6	6	6
<b>RV15</b>	7	7	7	7	7
<b>RV20</b>	8	8	8	8	8
<b>RV30</b>	9	9	10	9	9
<b>abs-r</b>	12	12	12	12	12
<b>hl</b>	11	11	11	11	11
<b>maxar2</b>	10	10	9	10	10
<b>RAV5</b>	3	3	3	3	3
<b>RAV10</b>	5	5	5	5	5
<b>RVHL10</b>	2	2	2	2	2
<b>RAVHL10</b>	1	1	1	1	1

# Empirical Analysis

Optimal weights  $\hat{w}$  :

Proxy	weight
RAV10HIGH	1.04
RAV10LOW	0.72
RAV10	-0.76

$H^{(\hat{w})}$  has  $PV = 0.038$

# Empirical Analysis

Stability of optimized proxy  $H^{(\hat{w})}$

	full	1st	2nd	3rd	4th
name	PV	PV	PV	PV	PV
$H^{(\hat{w})}$	0.038	0.039	0.043	0.043	0.028
$H^{(\hat{w},1)}$	0.038	0.039	0.043	0.043	0.028
$H^{(\hat{w},2)}$	0.039	0.039	0.043	0.043	0.029
$H^{(\hat{w},3)}$	0.039	0.040	0.044	0.042	0.029
$H^{(\hat{w},4)}$	0.039	0.040	0.045	0.044	0.027

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# Empirical Analysis

Heuristic: forecast comparison

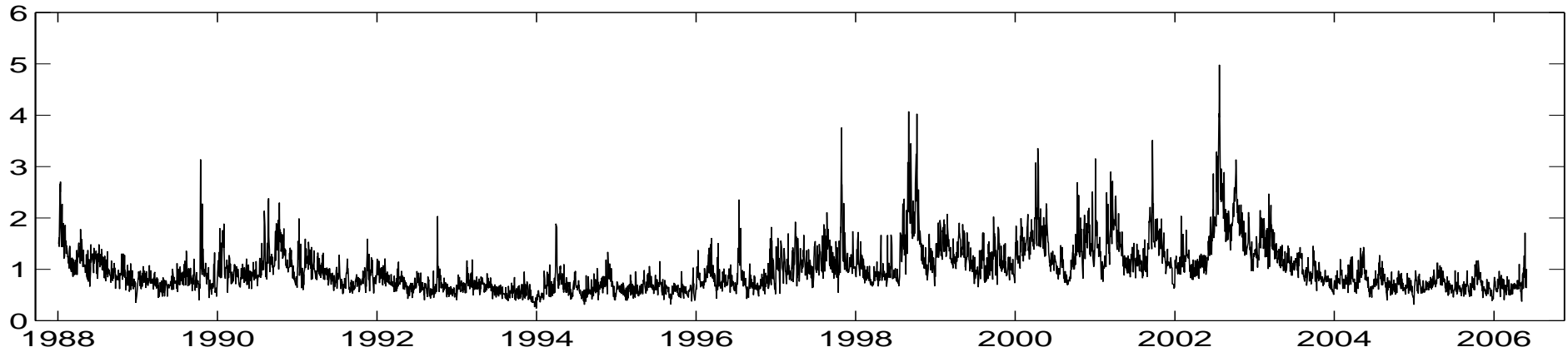
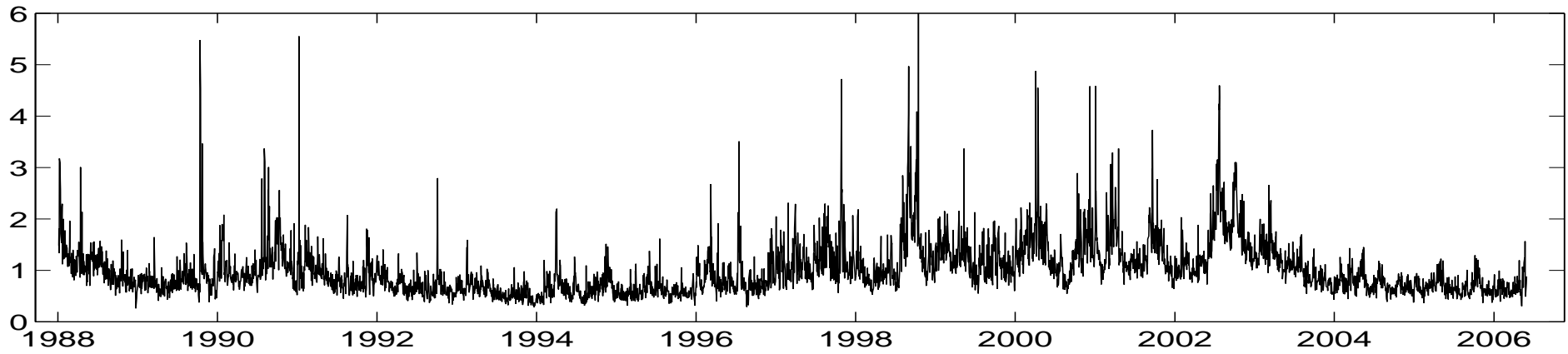
	RV30	RV20	RV15	RV10	RV5	$H^{(\hat{w})}$
RV30(-1)	0.35	0.39	0.42	0.46	0.50	0.58
RV20(-1)	0.38	0.42	0.45	0.49	0.54	0.61
RV15(-1)	0.39	0.44	0.47	0.50	0.55	0.63
RV10(-1)	0.41	0.45	0.48	0.52	0.57	0.66
RV5(-1)	0.43	0.48	0.51	0.55	0.60	0.69
$H^{(\hat{w})}(-1)$	0.44	0.48	0.51	0.54	0.60	0.71

$R^2$  of regression

$$\log(H^{(j)}(R_n)) = \alpha + \beta \log(H^{(i)}(R_{n-1})) + \varepsilon_n$$

# Empirical Analysis

RV5 vs  $H(\hat{w})$

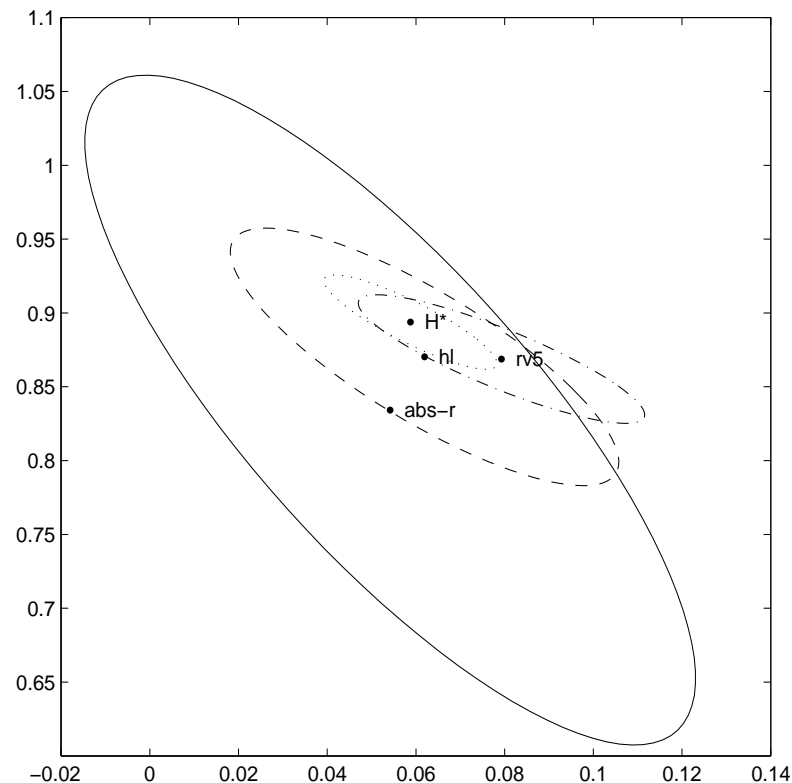


# Future Research

Parameter estimation: consider the Garch(1,1) recursion

$$s_n^2 = \kappa + \alpha r_{n-1}^2 + \beta s_{n-1}^2$$

Decreasing the confidence region in the  $(\alpha, \beta)$  plane:



Thank You!