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#### What is asymptotically testable and what is not?

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## Asymptotic symmetric testing

Observe *i.i.d.* data  $X^n \sim P^n$ , model  $P \in \mathscr{P}$ ; for disjoint  $B, V \subset \mathscr{P}$ ,

 $H_0$ :  $P \in B$ , or  $H_1$ :  $P \in V$ .

Look for test functions  $\phi_n : \mathscr{X}^n \to [0, 1]$  s.t.

 $P^n \phi_n(X^n) \to 0$ , and  $Q^n(1 - \phi_n(X^n)) \to 0$ 

for all  $P \in B$  and all  $Q \in V$ .

Equivalently, we want,

A testing procedure that chooses for B or V based on  $X^n \sim P^n$ for every  $n \ge 1$ , has property (D) if it is wrong only a finite number of times with  $P^{\infty}$ -probability one.

Property (D) is sometimes referred to as "discernibility".

### Some examples and unexpected answers (I)

Consider non-parametric regression with  $f : X \to \mathbb{R}$  and test for smoothness,

$$H_0: f \in C^1(X \to \mathbb{R}), \quad H_1: f \in C^2(X \to \mathbb{R}),$$

Consider a non-parametric density estimation with  $p : \mathbb{R} \to [0, \infty)$  and test for square-integrability,

$$H_0$$
:  $\int x^2 p(x) dx < \infty$ ,  $H_1$ :  $\int x^2 p(x) dx = \infty$ .

Practical problem we cannot use the data to determine with asymptotic certainty, if CLT applies with our data.

# Some examples and unexpected answers (II)

Coin-flip  $X^n \sim \text{Bernoulli}(p)^n$  with  $p \in [0, 1]$ .

Consider Cover's rational mean problem (1973):

 $H_0$ :  $p \in [0, 1] \cap \mathbb{Q}$ ,  $H_1$ :  $p \in [0, 1] \setminus \mathbb{Q}$ .

Consider also Dembo and Peres's irrational alternative (1995):

 $H_0: p \in [0,1] \cap \mathbb{Q}, \quad H_1: p \in [0,1] \cap \sqrt{2} + \mathbb{Q},$ 

Consider ultimately fractal hypotheses, e.g. with Cantor set C,

 $H_0$ :  $p \in C$ ,  $H_1$ :  $p \in [0, 1] \setminus C$ .

### Three forms of testability

**Definition 5.1**  $(\phi_n)$  is a uniform test sequence for B vs V, if,

$$\sup_{P \in B} P^n \phi_n \to 0, \quad \sup_{Q \in V} Q^n (1 - \phi_n) \to 0.$$
(1)

**Definition 5.2**  $(\phi_n)$  is a pointwise test sequence for B vs V, if,

$$\phi_n(X^n) \xrightarrow{P} 0, \quad \phi_n(X^n) \xrightarrow{Q} 1,$$
 (2)

for all  $P \in B$  and  $Q \in V$ .

**Definition 5.3**  $(\phi_n)$  is a Bayesian test sequence for B vs V, if,

$$\phi_n(X^n) \xrightarrow{P} 0, \quad \phi_n(X^n) \xrightarrow{Q} 1,$$
 (3)

for  $\Pi$ -almost-all  $P \in B$  and  $Q \in V$ .

### Posterior odds model selection for frequentists

Johnson & Rossell (JRSSB, 2010), Taylor & Tibshirani (PNAS, 2016)

**Theorem 6.1** Given measurable  $B, V \subset \Theta$  ( $\Pi(B), \Pi(V) > 0$ ) and,

i. there are Bayesian tests for B vs V of power  $a_n \downarrow 0$ ,

ii.

$$\int_{B} P^{n} \phi_{n} d\Pi(P) + \int_{V} Q^{n} (1 - \phi_{n}) d\Pi(Q) = o(a_{n}),$$
  
and, for all  $P \in B$ ,  $P^{n} \triangleleft a_{n}^{-1} P_{n}^{\Pi|B}$ ; for all  $Q \in V$ ,  $Q^{n} \triangleleft a_{n}^{-1} P_{n}^{\Pi|V}$ ,

then posterior odds give rise to a pointwise test for B vs V.

See BK, "The frequentist validity of Bayesian limits", arXiv:1611.08444 [math.ST]

#### Example: KL-neighbourhoods

**Definition 7.1** Given  $(P_n)$ ,  $(Q_n)$  and a  $a_n \downarrow 0$ ,  $Q_n$  is  $a_n$ -remotely contiguous w.r.t.  $P_n$   $(Q_n \triangleleft a_n^{-1}P_n)$ , if for any msb  $\psi_n : \mathscr{X}^n \to [0, 1]$ 

 $P_n\psi_n = o(a_n) \quad \Rightarrow \quad Q_n\psi_n = o(1)$ 

**Example 7.2** Let  $\mathscr{P}$  be a model for i.i.d. data  $X^n$ . Let  $P_0, P$  and  $\epsilon > 0$  be such that  $-P_0 \log(dP/dP_0) < \epsilon^2$ . Then, for large enough n,

$$\frac{dP^n}{dP_0^n}(X^n) \ge e^{-\frac{n}{2}\epsilon^2},\tag{4}$$

with  $P_0^n$ -probability one. So for any tests  $\psi_n$ ,

$$P^n \psi_n \ge e^{-\frac{1}{2}n\epsilon^2} P_0^n \psi_n. \tag{5}$$

So if  $P^n \phi_n = o(\exp(-\frac{1}{2}n\epsilon^2))$  then  $P_0^n \phi_n = o(1)$ :  $P_0^n \triangleleft a_n^{-1}P^n$  with  $a_n = \exp(-\frac{1}{2}n\epsilon^2)$ .

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# Example: select the DAG (I)

Observe an *i.i.d.*  $X^n$  of vectors of discrete random variables  $X_i = (X_{1,i}, \ldots, X_{k,i}) \in \mathbb{Z}^k$ ,  $1 \le i \le n$ .

Define a family  $\mathscr{F}$  of kernels  $p_{\theta}(\cdot|\cdot) : \mathbb{Z} \times \mathbb{Z}^{l} \to [0, 1]$ , for  $\theta \in \Theta$ ,  $1 \leq l \leq k$ . Assume that  $\Theta$  is compact and,

$$heta\mapsto \sum_{x\in\mathbb{Z}}f(x)P_{ heta}(x|z_1,\ldots,z_l)$$

is continuous, for every bounded  $f : \mathbb{Z} \to \mathbb{R}$  and all  $z_1, \ldots, z_l \in \mathbb{Z}$ .

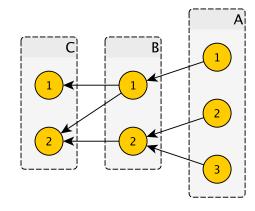
 $X \sim P$  follows a graphical model,

$$P_{\mathscr{A},\theta}(X_1 \in B_1, \ldots, X_k \in B_k) = \prod_{i=1}^k P_{\theta_i}(X_i \in B_i | \mathscr{A}_i)$$

where  $\mathscr{A}_i \subset \{1, \ldots, k\}$  denotes the *parents* of  $X_i$  (and  $\mathscr{A}_{ij} = \mathscr{A}_i \cup \mathscr{A}_j$ ). Together, the  $\mathscr{A}_i$  describe a directed, a-cyclical graph (DAG).

### Example: select the DAG (II)

The DAG  $\mathscr{A} = (\mathscr{A}_i : 1 \leq i \leq k)$  represents a number of conditional independence statements concerning the components  $X_1, \ldots, X_k$ .



$$P_{\mathscr{A},\theta}(C_{1} \in \cdot, \dots, A_{3} \in \cdot)$$
  
=  $P_{\theta_{C,1}}(\cdot|B_{1}) \times P_{\theta_{C,2}}(\cdot|B_{1}, B_{2})$   
 $\times P_{\theta_{B,1}}(\cdot|A_{1}) \times P_{\theta_{B_{2}}}(\cdot|A_{2}, A_{3})$   
 $\times P_{\theta_{A,1}}(\cdot) \times P_{\theta_{A,2}}(\cdot) \times P_{\theta_{A,3}}(\cdot)$ 

Fig 1. An small example DAG: No arrow means  $X_i \perp X_j | \mathscr{A}_{ij}$ .  $\mathscr{A}_{C_1} = \{B_1\}, \mathscr{A}_{B_2} = \{A_2, A_3\}$ , so given  $B_1, A_2$  and  $A_3, C_1$  is independent of  $B_2$ .

### Example: select the DAG (III)

Define the submodels  $\mathscr{P}_{\mathscr{A}} = \{P_{\mathscr{A},\theta} : \theta \in \Theta^k\}$ , for all  $\mathscr{A}$ . Given any  $\mathscr{A}' \neq \mathscr{A}$ , there is a pair  $X_i \perp X_j | \mathscr{A}_{ij}$  but  $X_i \not\perp X_j | \mathscr{A}'_{ij}$ .

Require that, for all  $\theta$ , all  $A, B \subset \mathbb{Z}$ ,

 $\left|P_{\mathscr{A}',\theta}(X_i \in A, X_j \in B|\mathscr{A}_{ij}) - P_{\mathscr{A}',\theta}(X_i \in A|\mathscr{A}_{ij}) P_{\mathscr{A}',\theta}(X_j \in B|\mathscr{A}_{ij})\right| > \epsilon,$ for some  $\epsilon > 0$  that depends only on  $\mathscr{A}$  and  $\mathscr{A}'$ .

With a KL-prior posterior odds for  $\mathscr{P}_{\mathscr{A}}$  select the correct DAG  $\mathscr{A}$ .

## Uniform testability: equivalent formulations

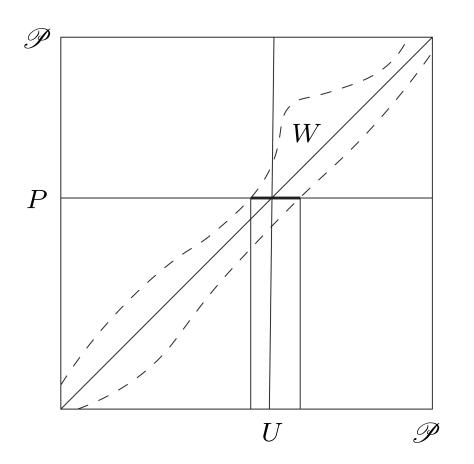
**Proposition 11.1** Let  $\mathscr{P}$  be a model for *i.i.d.* data with disjoint B and V. The following are equivalent:

i. there exists a uniform test sequence  $(\phi_n)$ ,

$$\sup_{P\in B}P^n\phi_n
ightarrow 0,\quad \sup_{Q\in V}Q^n(1-\phi_n)
ightarrow 0,$$

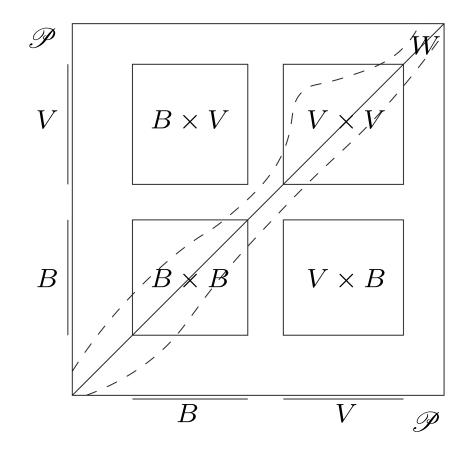
ii. there is a exponentially powerful uniform test sequence  $(\psi_n)$ ,

$$\sup_{P \in B} P^n \psi_n \leq e^{-nD}, \quad \sup_{Q \in V} Q^n (1 - \psi_n) \leq e^{-nD}$$



**Fig 2.** Let  $P \in \mathscr{P}$  and entourage  $W \in \mathscr{U}_{\infty}$  be given. Define neighbourhood  $U \in \mathscr{T}_{\infty}$  as  $U = \{Q \in \mathscr{P} : (Q, P) \in W\}$ 

# Uniform separation (II)



**Fig 3.** *B* and *V* are uniformly separated by  $\mathscr{U}_{\infty}$  if there is a  $W \in \mathscr{U}_{\infty}$  that does not meet  $B \times V$  and  $V \times B$ .

#### Characterisation of uniform testability

**Theorem 14.1** Let  $\mathscr{P}$  be a model for *i.i.d.* data with disjoint B and V. The following are equivalent:

(i.) there are uniform tests  $\phi_n$  for B versus V,

(ii.) B and V are uniformly separated by  $\mathscr{U}_{\infty}$ .

**Corollary 14.2** (Parametrised models) Suppose  $\mathscr{P} = \{p_{\theta} : \theta \in \Theta\}$ , with  $(\Theta, d)$  compact, metric space and  $\theta \to P_{\theta}$  identifiable and  $\mathscr{T}_{\infty}$ continuous, (that is, for every  $f \in \mathscr{F}_n$ ,  $\theta \mapsto \int f dP_{\theta}^n$  is continuous). If  $B_0, V_0 \subset \Theta$  with  $d(B_0, V_0) > 0$ , then the images  $B = \{P_{\theta} : \theta \in B_0\}$ ,  $V = \{P_{\theta} : \theta \in V_0\}$  are uniformly testable.

#### Pointwise testability: equivalent formulations

**Proposition 15.1** Let  $\mathscr{P}$  be a model for i.i.d. data and let B, V be disjoint model subsets. The following are equivalent:

i. there are tests  $(\phi_n)$  such that, for all  $P \in B$  and  $Q \in V$ ,

 $P^n\phi_n 
ightarrow 0, \quad Q^n(1-\phi_n)
ightarrow 0,$ 

ii. there are tests  $(\phi_n)$  such that, for all  $P \in B$  and  $Q \in V$ ,

$$\phi_n(X^n) \xrightarrow{P} 0, \quad (1 - \phi_n(X^n)) \xrightarrow{Q} 0,$$

iii. there are tests  $(\phi_n)$  such that, for all  $P \in B$  and  $Q \in V$ ,

$$\phi_n(X^n) \xrightarrow{P-a.s.} 0, \quad (1 - \phi_n(X^n)) \xrightarrow{Q-a.s.} 0$$

### Pointwise testability in dominated models

**Definition 16.1** The testing problem has a (uniform) representation on X, if there exists a  $\mathscr{T}_{\infty}$ -(uniformly-)continuous, surjective map  $f: B \cup V \to X$  such that  $f(B) \cap f(V) = \emptyset$ .

**Definition 16.2** The model is parametrised by  $\Theta$ , if there exists a  $\mathscr{T}_{\infty}$ -continuous bijection  $P : \Theta \to \mathscr{P}$  (i.e. for every  $m \ge 1$  and measurable  $f : \mathscr{X}^m \to [0, 1]$ , the map  $\theta \mapsto \int f \, dP^m_{\theta}$  is continuous).

### Characterisation of pointwise testability

**Theorem 17.1** Let  $\mathscr{P}$  be a dominated model for *i.i.d.* data with disjoint B, V. The following are equivalent,

- i. there exists a pointwise test for B vs V,
- ii. the problem has a representation  $f : B \cup V \to X$  on a normal space X and there exist disjoint  $F_{\sigma}$ -sets  $B', V' \subset X$  such that  $f(B) \subset B', f(V) \subset V'$ ,
- iii. the problem has a uniform representation  $\psi : B \cup V \to X$  on a separable, metrizable space X with  $\psi(B), \psi(V)$  both  $F_{\sigma}$  and  $G_{\delta}$ -sets.

# Finite entropy and uniform integrability

**Corollary 18.1** Suppose that  $\mathscr{P}$  is dominated and  $\mathsf{TV}$ -totally-bounded. Then disjoint  $B, V \subset \mathscr{P}$  are pointwise testable, if and only if, B, V are both  $F_{\sigma}$ - and  $G_{\delta}$ -sets in  $B \cup V$  (for  $\mathscr{T}_{TV}$ ).

**Corollary 18.2** Suppose that  $\mathscr{P}$  is dominated by a probability measure, with a uniformly integrable family of densities. Then disjoint  $B, V \subset \mathscr{P}$  are pointwise testable, if and only if, B, V are both  $F_{\sigma}$ - and  $G_{\delta}$ -sets in  $B \cup V$  (for  $\mathscr{T}_{C}$ ).

### Bayesian testability: equivalent formulations

**Theorem 19.1** Let a model  $(\mathscr{P}, \mathscr{G}, \Pi)$  with  $B, V \in \mathscr{G}$  be given, with  $\Pi(B) > 0, \Pi(V) > 0$ . The following are equivalent,

- i. there exist Bayesian tests for B vs V,
- ii. there are tests  $\phi_n$  such that for  $\Pi$ -almost-all  $P \in B, Q \in V$ ,

$$P^n \phi_n \to 0, \quad Q^n (1 - \phi_n) \to 0,$$

iii. there are tests  $\phi_n : \mathscr{X}^n \to [0, 1]$  such that,

$$\int_{B} P^{n} \phi_{n} d\Pi(P) + \int_{V} Q^{n} (1 - \phi_{n}) d\Pi(Q) \to 0,$$

iv. for  $\Pi$ -almost-all  $P \in B$ ,  $Q \in V$ ,

$$\Pi(V|X^n) \xrightarrow{P} 0, \quad \Pi(B|X^n) \xrightarrow{Q} 0.$$

# Characterisation of Bayesian testability

**Theorem 20.1** Let  $(\mathscr{P}, \mathscr{G})$  be a measurable model with a prior  $\Pi$  that is a Radon measure and hypotheses B, V. There is a Bayesian test sequence for B vs V, if and only if, B, V are  $\mathscr{G}$ -measurable.

### Consistent model selection

Let  $\mathscr{P}$  be a model for *i.i.d.* data  $X^n \sim P^n$ ,  $(n \ge 1)$ , and suppose that  $(\mathscr{P}, \mathscr{G}, \Pi)$  has finite, measurable partition,

 $P \in \mathscr{P} = \mathscr{P}_1 \cup \ldots \cup \mathscr{P}_M.$ 

Model-selection Which  $1 \leq i \leq M$ ? (such that  $P \in \mathscr{P}_i$ )

**Theorem 21.1** Assume that for all  $1 \le i < j \le M$ ,

 $\mathscr{P}_i$  and  $\mathscr{P}_i$  are  $\mathscr{U}_{\infty}$ -uniformly separated.

Let  $1 \leq i \leq M$  be such that  $P \in \mathscr{P}_i$ . If  $\Pi$  is a KL-prior, then indicators for posterior odds,

$$\phi_n(X^n) = \mathbb{1}\Big\{X^n : \Pi(\mathscr{P}_i | X^n) \ge \sum_{j \neq i} \Pi(\mathscr{P}_j | X^n)\Big\},\$$

are a pointwise test for  $\mathscr{P}_i$  vs  $\cup_{j\neq i} \mathscr{P}_j$ .

# Thank you for your attention

BK, "The frequentist validity of Bayesian limits" arXiv:1611.08444 [math.ST]

#### Remote contiguity

**Definition 23.1** Given  $(P_n)$ ,  $(Q_n)$  and a  $a_n \downarrow 0$ ,  $Q_n$  is  $a_n$ -remotely contiguous w.r.t.  $P_n$   $(Q_n \triangleleft a_n^{-1}P_n)$ , if for any msb  $\psi_n : \mathscr{X}^n \to [0, 1]$ 

$$P_n\psi_n = o(a_n) \quad \Rightarrow \quad Q_n\psi_n = o(1)$$

**Lemma 23.2**  $Q_n \triangleleft a_n^{-1}P_n$  if any of the following holds:

(i) For any bnd msb  $T_n : \mathscr{X}^n \to \mathbb{R}, a_n^{-1}T_n \xrightarrow{P_n} 0$ , implies  $T_n \xrightarrow{Q_n} 0$ 

(ii) Given  $\epsilon > 0$ , there is a  $\delta > 0$  s.t.  $Q_n(dP_n/dQ_n < \delta a_n) < \epsilon$  f.l.e.n.

(iii) There is a b > 0 s.t.  $\liminf_{n \to \infty} b a_n^{-1} P_n(dQ_n/dP_n > b a_n^{-1}) = 1$ 

(iv) Given  $\epsilon > 0$ , there is a c > 0 such that  $\|Q_n - Q_n \wedge c a_n^{-1} P_n\| < \epsilon$ 

(v) Under  $Q_n$ , every subsequence of  $(a_n(dP_n/dQ_n)^{-1})$  has a further subsequence that converges in  $\mathcal{T}_C$ .

### The model as a uniform space

Take  $\mathscr{X}$  a separable metrizable space, with Borel  $\sigma$ -algebra  $\mathscr{B}$ .

The class  $\mathscr{F}_n$  contains all bounded,  $\mathscr{B}^n$ -measurable  $f : \mathscr{X}^n \to \mathbb{R}$ .

For every  $n \ge 1$  and  $f \in \mathscr{F}_n$ , define the entourage,

$$W_{n,f} = \{ (P,Q) \in \mathscr{P} \times \mathscr{P} : |P^n f - Q^n f| < 1 \}.$$

Defines uniformity  $\mathscr{U}_n$  (with topology  $\mathscr{T}_n$ ). Take  $\mathscr{U}_{\infty} = \bigcup_{n>1} \mathscr{U}_n$ .

$$P \to Q \text{ in } \mathscr{T}_{\infty} \quad \Leftrightarrow \quad \int f \, dP^n \to \int f \, dQ^n,$$

for all  $n \ge 1$  and all  $f \in \mathscr{F}_n$ . Note also,

$$\mathscr{U}_C \subset \mathscr{U}_1 \subset \cdots \subset \mathscr{U}_\infty \subset \mathscr{U}_{TV}.$$

# The Dunford-Pettis theorem

**Theorem 25.1** (Dunford-Pettis) Assume  $\mathscr{P}$  is dominated by a probability measure Q with densities in  $\mathscr{P}_Q \subset L^1(Q)$ ;  $\mathscr{P}_Q$  is relatively weakly compact, if and only if, for every  $\epsilon > 0$  there is an M > 0 such that,

$$\sup_{P\in\mathscr{P}}\int_{\{dP/dQ>M\}}\frac{dP}{dQ}\,dQ<\epsilon,$$

that is,  $\mathscr{P}_Q$  is uniformly Q-integrable.

### Uniform separation

**Definition 26.1** Subsets  $B, V \subset \mathscr{P}$  are uniformly separated by  $\mathscr{U}_{\infty}$ , if there exists an entourage  $W \in \mathscr{U}_{\infty}$  such that,

 $(B \times V \cup V \times B) \cap W = \emptyset.$ 

In other words, there are  $J, m \ge 1$ ,  $\epsilon > 0$  and bounded, measurable functions  $f_1, \ldots, f_J : \mathscr{X}^m \to [0, 1]$  such that, for any  $P, Q \in B \cup V$ , if,

$$\max_{1 \le j \le J} \left| P^m f_j - Q^m f_j \right| < \epsilon,$$

then either  $P, Q \in B$ , or  $P, Q \in V$ . (If the model is  $\mathcal{T}_{\infty}$ -compact, m = 1 suffices).

# The Le Cam-Schwartz theorem

**Theorem 27.1** (Le Cam-Schwartz, 1960) Let  $\mathscr{P}$  be a model for *i.i.d.* data  $X^n$  with disjoint subsets B, V. The following are equivalent:

- i. there exist (uniformly) consistent tests for B vs V,
- ii. there is a sequence of  $\mathscr{U}_{\infty}$ -uniformly continuous  $\psi_n : \mathscr{P} \to [0,1]$ ,

$$\psi_n(P) \to \mathbf{1}_V(P),$$
 (6)

(uniformly) for all  $P \in B \cup V$ .

### Example: how many clusters? (I)

Observe *i.i.d.*  $X^n \sim P^n$ , where P dominated with density p.

Clusters Family  $\mathscr{F}$  of kernels  $\varphi_{\theta} : \mathbb{R} \to [0, \infty)$ , with parameter  $\theta \in \Theta$ . Assume  $\Theta$  compact and,

 $\theta \mapsto \int f(x)\varphi_{\theta}(x) \, dx,$ 

is continuous, for every bounded, measurable  $f:\mathbb{R}\to\mathbb{R}.$  Define  $\Theta_M'=\Theta^M/\sim.$ 

Model Assume that there is an M > 0 such that p can be written as,

$$p_{\lambda,\theta}(x) = \sum_{m=1}^{M} \lambda_m p_{\theta_m}(x),$$

for some  $M \ge 1$ , with  $\lambda \in S_M = \{\lambda \in [0,1]^M : \sum_m \lambda_m = 1\}$ ,  $\theta \in \Theta'_M$ .

# Example: how many clusters? (II)

Assume *M* less than some known *M'*. Choose prior  $\Pi_{\lambda,M}$  for  $\lambda \in S_M$  such that, for some  $\epsilon > 0$ ,

$$\Pi_{\lambda,M} (\lambda \in S_M : \epsilon < \min\{\lambda_m\}, \max\{\lambda_m\} < 1 - \epsilon) = 1$$

For  $\theta \in \Theta'_M$  also choose a prior  $\Pi_{\theta,M}$  that 'stays away from the edges'. Define,

$$\Pi = \sum_{M=1}^{M'} \mu_M \Pi_{\lambda,M} \times \Pi_{\theta,M}.$$

(for  $\sum_M \mu_M = 1$ ).

If  $\Pi$  is a KL-prior, posterior odds select the correct number of clusters M. If there are no M' and  $\epsilon$  known, there are sequences  $M'_n \to \infty$  and  $\epsilon_n \downarrow 0$  with priors  $\Pi_n$  that finds the correct number of clusters.