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# What is asymptotically testable and what is not? 

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## Asymptotic symmetric testing

Observe i.i.d. data $X^{n} \sim P^{n}$, model $P \in \mathscr{P}$; for disjoint $B, V \subset \mathscr{P}$,

$$
H_{0}: P \in B, \quad \text { or } \quad H_{1}: P \in V .
$$

Look for test functions $\phi_{n}: \mathscr{X}^{n} \rightarrow[0,1]$ s.t.

$$
P^{n} \phi_{n}\left(X^{n}\right) \rightarrow 0, \quad \text { and } \quad Q^{n}\left(1-\phi_{n}\left(X^{n}\right)\right) \rightarrow 0
$$

for all $P \in B$ and all $Q \in V$.

Equivalently, we want,

A testing procedure that chooses for $B$ or $V$ based on $X^{n} \sim P^{n}$ for every $n \geq 1$, has property ( $D$ ) if it is wrong only a finite number of times with $P^{\infty}$-probability one.

Property (D) is sometimes referred to as "discernibility".

## Some examples and unexpected answers (I)

Consider non-parametric regression with $f: X \rightarrow \mathbb{R}$ and test for smoothness,

$$
H_{0}: f \in C^{1}(X \rightarrow \mathbb{R}), \quad H_{1}: f \in C^{2}(X \rightarrow \mathbb{R})
$$

Consider a non-parametric density estimation with $p: \mathbb{R} \rightarrow[0, \infty)$ and test for square-integrability,

$$
H_{0}: \int x^{2} p(x) d x<\infty, \quad H_{1}: \int x^{2} p(x) d x=\infty
$$

Practical problem we cannot use the data to determine with asymptotic certainty, if CLT applies with our data.

## Some examples and unexpected answers (II)

Coin-flip $X^{n} \sim \operatorname{Bernoulli}(p)^{n}$ with $p \in[0,1]$.

Consider Cover's rational mean problem (1973):

$$
H_{0}: p \in[0,1] \cap \mathbb{Q}, \quad H_{1}: p \in[0,1] \backslash \mathbb{Q} .
$$

Consider also Dembo and Peres's irrational alternative (1995):

$$
H_{0}: p \in[0,1] \cap \mathbb{Q}, \quad H_{1}: p \in[0,1] \cap \sqrt{2}+\mathbb{Q}
$$

Consider ultimately fractal hypotheses, e.g. with Cantor set $C$,

$$
H_{0}: p \in C, \quad H_{1}: p \in[0,1] \backslash C
$$

## Three forms of testability

Definition 5.1 ( $\phi_{n}$ ) is a uniform test sequence for $B$ vs $V$, if,

$$
\begin{equation*}
\sup _{P \in B} P^{n} \phi_{n} \rightarrow 0, \quad \sup _{Q \in V} Q^{n}\left(1-\phi_{n}\right) \rightarrow 0 . \tag{1}
\end{equation*}
$$

Definition $5.2\left(\phi_{n}\right)$ is a pointwise test sequence for $B$ vs $V$, if,

$$
\begin{equation*}
\phi_{n}\left(X^{n}\right) \xrightarrow{P} 0, \quad \phi_{n}\left(X^{n}\right) \xrightarrow{Q} 1, \tag{2}
\end{equation*}
$$

for all $P \in B$ and $Q \in V$.
Definition 5.3 ( $\phi_{n}$ ) is a Bayesian test sequence for $B$ vs $V$, if,

$$
\begin{equation*}
\phi_{n}\left(X^{n}\right) \xrightarrow{P} 0, \quad \phi_{n}\left(X^{n}\right) \xrightarrow{Q} 1, \tag{3}
\end{equation*}
$$

for $\Pi$-almost-all $P \in B$ and $Q \in V$.

## Posterior odds model selection for frequentists

## Johnson \& Rossell (JRSSB, 2010), Taylor \& Tibshirani (PNAS, 2016)

Theorem 6.1 Given measurable $B, V \subset \Theta(\Pi(B), \Pi(V)>0)$ and,
i. there are Bayesian tests for $B$ vs $V$ of power $a_{n} \downarrow 0$,

$$
\int_{B} P^{n} \phi_{n} d \Pi(P)+\int_{V} Q^{n}\left(1-\phi_{n}\right) d \Pi(Q)=o\left(a_{n}\right)
$$

ii. and, for all $P \in B, P^{n} \triangleleft a_{n}^{-1} P_{n}^{\Pi \mid B}$; for all $Q \in V, Q^{n} \triangleleft a_{n}^{-1} P_{n}^{\Pi \mid V}$,
then posterior odds give rise to a pointwise test for $B$ vs $V$.

See BK, "The frequentist validity of Bayesian limits", arXiv:1611.08444 [math.ST]

## Example: KL-neighbourhoods

Definition 7.1 Given $\left(P_{n}\right),\left(Q_{n}\right)$ and $a a_{n} \downarrow 0, Q_{n}$ is $a_{n}$-remotely contiguous w.r.t. $P_{n}\left(Q_{n} \triangleleft a_{n}^{-1} P_{n}\right)$, if for any $m s b \psi_{n}: \mathscr{X}^{n} \rightarrow[0,1]$

$$
P_{n} \psi_{n}=o\left(a_{n}\right) \quad \Rightarrow \quad Q_{n} \psi_{n}=o(1)
$$

Example 7.2 Let $\mathscr{P}$ be a model for i.i.d. data $X^{n}$. Let $P_{0}, P$ and $\epsilon>0$ be such that $-P_{0} \log \left(d P / d P_{0}\right)<\epsilon^{2}$. Then, for large enough $n$,

$$
\begin{equation*}
\frac{d P^{n}}{d P_{0}^{n}}\left(X^{n}\right) \geq e^{-\frac{n}{2} \epsilon^{2}} \tag{4}
\end{equation*}
$$

with $P_{0}^{n}$-probability one. So for any tests $\psi_{n}$,

$$
\begin{equation*}
P^{n} \psi_{n} \geq e^{-\frac{1}{2} n \epsilon^{2}} P_{0}^{n} \psi_{n} \tag{5}
\end{equation*}
$$

So if $P^{n} \phi_{n}=o\left(\exp \left(-\frac{1}{2} n \epsilon^{2}\right)\right)$ then $P_{0}^{n} \phi_{n}=o(1)$ : $P_{0}^{n} \triangleleft a_{n}^{-1} P^{n}$ with $a_{n}=\exp \left(-\frac{1}{2} n \epsilon^{2}\right)$.

## Example: select the DAG (I)

Observe an i.i.d. $X^{n}$ of vectors of discrete random variables $X_{i}=$ $\left(X_{1, i}, \ldots, X_{k, i}\right) \in \mathbb{Z}^{k}, 1 \leq i \leq n$.

Define a family $\mathscr{F}$ of kernels $p_{\theta}(\cdot \mid \cdot): \mathbb{Z} \times \mathbb{Z}^{l} \rightarrow[0,1]$, for $\theta \in \Theta, 1 \leq l \leq k$. Assume that $\Theta$ is compact and,

$$
\theta \mapsto \sum_{x \in \mathbb{Z}} f(x) P_{\theta}\left(x \mid z_{1}, \ldots, z_{l}\right)
$$

is continuous, for every bounded $f: \mathbb{Z} \rightarrow \mathbb{R}$ and all $z_{1}, \ldots, z_{l} \in \mathbb{Z}$.
$X \sim P$ follows a graphical model,

$$
P_{\mathscr{A}, \theta}\left(X_{1} \in B_{1}, \ldots, X_{k} \in B_{k}\right)=\prod_{i=1}^{k} P_{\theta_{i}}\left(X_{i} \in B_{i} \mid \mathscr{A}_{i}\right)
$$

where $\mathscr{A}_{i} \subset\{1, \ldots, k\}$ denotes the parents of $X_{i}\left(\right.$ and $\left.\mathscr{A}_{i j}=\mathscr{A}_{i} \cup \mathscr{A}_{j}\right)$. Together, the $\mathscr{A}_{i}$ describe a directed, a-cyclical graph (DAG).

## Example: select the DAG (II)

The $\operatorname{DAG} \mathscr{A}=\left(\mathscr{A}_{i}: 1 \leq i \leq k\right)$ represents a number of conditional independence statements concerning the components $X_{1}, \ldots, X_{k}$.


$$
\begin{aligned}
& P_{\mathscr{A}, \theta}\left(C_{1} \in \cdot, \ldots, A_{3} \in \cdot\right) \\
& =P_{\theta_{C, 1}}\left(\cdot \mid B_{1}\right) \times P_{\theta_{C, 2}}\left(\cdot \mid B_{1}, B_{2}\right) \\
& \quad \times P_{\theta_{B, 1}}\left(\cdot \mid A_{1}\right) \times P_{\theta_{B_{2}}}\left(\cdot \mid A_{2}, A_{3}\right) \\
& \quad \times P_{\theta_{A, 1}}(\cdot) \times P_{\theta_{A, 2}}(\cdot) \times P_{\theta_{A, 3}}(\cdot)
\end{aligned}
$$

Fig 1. An small example DAG: No arrow means $X_{i} \perp X_{j} \mid \mathscr{A}_{i j}$. $\mathscr{A}_{C_{1}}=$ $\left\{B_{1}\right\}, \mathscr{A}_{B_{2}}=\left\{A_{2}, A_{3}\right\}$, so given $B_{1}, A_{2}$ and $A_{3}, C_{1}$ is independent of $B_{2}$.

## Example: select the DAG (III)

Define the submodels $\mathscr{P}_{\mathscr{A}}=\left\{P_{\mathscr{A}, \theta}: \theta \in \Theta^{k}\right\}$, for all $\mathscr{A}$. Given any $\mathscr{A}^{\prime} \neq \mathscr{A}$, there is a pair $X_{i} \perp X_{j} \mid \mathscr{A}_{i j}$ but $X_{i} \not \perp X_{j} \mid \mathscr{A}_{i j}^{\prime}$.

Require that, for all $\theta$, all $A, B \subset \mathbb{Z}$,

$$
\left|P_{\mathscr{A}^{\prime}, \theta}\left(X_{i} \in A, X_{j} \in B \mid \mathscr{A}_{i j}\right)-P_{\mathscr{A}^{\prime}, \theta}\left(X_{i} \in A \mid \mathscr{A}_{i j}\right) P_{\mathscr{A}^{\prime}, \theta}\left(X_{j} \in B \mid \mathscr{A}_{i j}\right)\right|>\epsilon,
$$

for some $\epsilon>0$ that depends only on $\mathscr{A}$ and $\mathscr{A}^{\prime}$.

With a KL-prior posterior odds for $\mathscr{P}_{\mathscr{A}}$ select the correct DAG $\mathscr{A}$.

## Uniform testability: equivalent formulations

## Proposition 11.1 Let $\mathscr{P}$ be a model for i.i.d. data with disjoint $B$

 and $V$. The following are equivalent:i. there exists a uniform test sequence $\left(\phi_{n}\right)$,

$$
\sup _{P \in B} P^{n} \phi_{n} \rightarrow 0, \quad \sup _{Q \in V} Q^{n}\left(1-\phi_{n}\right) \rightarrow 0
$$

ii. there is a exponentially powerful uniform test sequence $\left(\psi_{n}\right)$,

$$
\sup _{P \in B} P^{n} \psi_{n} \leq e^{-n D}, \quad \sup _{Q \in V} Q^{n}\left(1-\psi_{n}\right) \leq e^{-n D}
$$

## The model as a uniform space



Fig 2. Let $P \in \mathscr{P}$ and entourage $W \in \mathscr{U}_{\infty}$ be given. Define neighbourhood $U \in \mathscr{T}_{\infty}$ as $U=\{Q \in \mathscr{P}:(Q, P) \in W\}$

## Uniform separation (II)



Fig 3. $B$ and $V$ are uniformly separated by $\mathscr{U}_{\infty}$ if there is a $W \in \mathscr{U}_{\infty}$ that does not meet $B \times V$ and $V \times B$.

## Characterisation of uniform testability

Theorem 14.1 Let $\mathscr{P}$ be a model for i.i.d. data with disjoint $B$ and
$V$. The following are equivalent:
(i.) there are uniform tests $\phi_{n}$ for $B$ versus $V$,
(ii.) $B$ and $V$ are uniformly separated by $\mathscr{U}_{\infty}$.

Corollary 14.2 (Parametrised models) Suppose $\mathscr{P}=\left\{p_{\theta}: \theta \in \Theta\right\}$, with $(\Theta, d)$ compact, metric space and $\theta \rightarrow P_{\theta}$ identifiable and $\mathscr{T}_{\infty^{-}}$ continuous, (that is, for every $f \in \mathscr{F}_{n}, \theta \mapsto \int f d P_{\theta}^{n}$ is continuous). If $B_{0}, V_{0} \subset \Theta$ with $d\left(B_{0}, V_{0}\right)>0$, then the images $B=\left\{P_{\theta}: \theta \in B_{0}\right\}$, $V=\left\{P_{\theta}: \theta \in V_{0}\right\}$ are uniformly testable.

## Pointwise testability: equivalent formulations

Proposition 15.1 Let $\mathscr{P}$ be a model for i.i.d. data and let $B, V$ be disjoint model subsets. The following are equivalent:
$i$. there are tests $\left(\phi_{n}\right)$ such that, for all $P \in B$ and $Q \in V$,

$$
P^{n} \phi_{n} \rightarrow 0, \quad Q^{n}\left(1-\phi_{n}\right) \rightarrow 0,
$$

ii. there are tests $\left(\phi_{n}\right)$ such that, for all $P \in B$ and $Q \in V$,

$$
\phi_{n}\left(X^{n}\right) \xrightarrow{P} 0, \quad\left(1-\phi_{n}\left(X^{n}\right)\right) \xrightarrow{Q} 0,
$$

iii. there are tests $\left(\phi_{n}\right)$ such that, for all $P \in B$ and $Q \in V$,

$$
\phi_{n}\left(X^{n}\right) \xrightarrow{P \text {-a.s. }} 0, \quad\left(1-\phi_{n}\left(X^{n}\right)\right) \xrightarrow{Q \text {-a.s. }} 0 .
$$

## Pointwise testability in dominated models

Definition 16.1 The testing problem has a (uniform) representation on $X$, if there exists a $\mathscr{T}_{\infty}$-(uniformly-)continuous, surjective map $f: B \cup V \rightarrow X$ such that $f(B) \cap f(V)=\varnothing$.

Definition 16.2 The model is parametrised by $\Theta$, if there exists a $\mathscr{T}_{\infty}$-continuous bijection $P$. : $\Theta \rightarrow \mathscr{P}$ (i.e. for every $m \geq 1$ and measurable $f: \mathscr{X}^{m} \rightarrow[0,1]$, the $\operatorname{map} \theta \mapsto \int f d P_{\theta}^{m}$ is continuous).

## Characterisation of pointwise testability

Theorem 17.1 Let $\mathscr{P}$ be a dominated model for i.i.d. data with disjoint $B, V$. The following are equivalent,
i. there exists a pointwise test for $B$ vs $V$,
ii. the problem has a representation $f: B \cup V \rightarrow X$ on a normal space $X$ and there exist disjoint $F_{\sigma}$-sets $B^{\prime}, V^{\prime} \subset X$ such that $f(B) \subset B^{\prime}, f(V) \subset V^{\prime}$,
iii. the problem has a uniform representation $\psi: B \cup V \rightarrow X$ on a separable, metrizable space $X$ with $\psi(B), \psi(V)$ both $F_{\sigma^{-}}$and $G_{\delta^{-}}$ sets.

## Finite entropy and uniform integrability

Corollary 18.1 Suppose that $\mathscr{P}$ is dominated and $T V$-totally-bounded. Then disjoint $B, V \subset \mathscr{P}$ are pointwise testable, if and only if, $B, V$ are both $F_{\sigma^{-}}$and $G_{\delta^{-}}$sets in $B \cup V\left(\right.$ for $\left.\mathscr{T}_{T V}\right)$.

Corollary 18.2 Suppose that $\mathscr{P}$ is dominated by a probability measure, with a uniformly integrable family of densities. Then disjoint $B, V \subset \mathscr{P}$ are pointwise testable, if and only if, $B, V$ are both $F_{\sigma^{-}}$and $G_{\delta}$-sets in $B \cup V\left(\right.$ for $\left.\mathscr{T}_{C}\right)$.

## Bayesian testability: equivalent formulations

Theorem 19.1 Let a model $(\mathscr{P}, \mathscr{G}, \Pi)$ with $B, V \in \mathscr{G}$ be given, with $\Pi(B)>0, \Pi(V)>0$. The following are equivalent,
i. there exist Bayesian tests for $B$ vs $V$,
ii. there are tests $\phi_{n}$ such that for $\Pi$-almost-all $P \in B, Q \in V$,

$$
P^{n} \phi_{n} \rightarrow 0, \quad Q^{n}\left(1-\phi_{n}\right) \rightarrow 0
$$

iii. there are tests $\phi_{n}: \mathscr{X}^{n} \rightarrow[0,1]$ such that,

$$
\int_{B} P^{n} \phi_{n} d \Pi(P)+\int_{V} Q^{n}\left(1-\phi_{n}\right) d \Pi(Q) \rightarrow 0
$$

iv. for $\Pi$-almost-all $P \in B, Q \in V$,

$$
\Pi\left(V \mid X^{n}\right) \xrightarrow{P} 0, \quad \Pi\left(B \mid X^{n}\right) \xrightarrow{Q} 0 .
$$

## Characterisation of Bayesian testability

Theorem 20.1 Let $(\mathscr{P}, \mathscr{G})$ be a measurable model with a prior $\Pi$ that is a Radon measure and hypotheses $B, V$. There is a Bayesian test sequence for $B$ vs $V$, if and only if, $B, V$ are $\mathscr{G}$-measurable.

## Consistent model selection

Let $\mathscr{P}$ be a model for i.i.d. data $X^{n} \sim P^{n},(n \geq 1)$, and suppose that $(\mathscr{P}, \mathscr{G}, \Pi)$ has finite, measurable partition,

$$
P \in \mathscr{P}=\mathscr{P}_{1} \cup \ldots \cup \mathscr{P}_{M}
$$

Model-selection Which $1 \leq i \leq M ?\left(\right.$ such that $\left.P \in \mathscr{P}_{i}\right)$
Theorem 21.1 Assume that for all $1 \leq i<j \leq M$,

$$
\mathscr{P}_{i} \text { and } \mathscr{P}_{j} \text { are } \mathscr{U}_{\infty} \text {-uniformly separated. }
$$

Let $1 \leq i \leq M$ be such that $P \in \mathscr{P}_{i}$. If $\Pi$ is a KL-prior, then indicators for posterior odds,

$$
\phi_{n}\left(X^{n}\right)=1\left\{X^{n}: \Pi\left(\mathscr{P}_{i} \mid X^{n}\right) \geq \sum_{j \neq i} \Pi\left(\mathscr{P}_{j} \mid X^{n}\right)\right\}
$$

are a pointwise test for $\mathscr{P}_{i}$ vs $\cup_{j \neq i} \mathscr{P}_{j}$.

# Thank you for your attention 

BK, "The frequentist validity of Bayesian limits" arXiv:1611.08444 [math.ST]

## Remote contiguity

Definition 23.1 Given $\left(P_{n}\right),\left(Q_{n}\right)$ and $a a_{n} \downarrow 0, Q_{n}$ is $a_{n}$-remotely contiguous w.r.t. $P_{n}\left(Q_{n} \triangleleft a_{n}^{-1} P_{n}\right)$, if for any $m s b \psi_{n}: \mathscr{X}^{n} \rightarrow[0,1]$

$$
P_{n} \psi_{n}=o\left(a_{n}\right) \quad \Rightarrow \quad Q_{n} \psi_{n}=o(1)
$$

Lemma 23.2 $Q_{n} \triangleleft a_{n}^{-1} P_{n}$ if any of the following holds:
(i) For any bnd msb $T_{n}: \mathscr{X}^{n} \rightarrow \mathbb{R}, a_{n}^{-1} T_{n} \xrightarrow{P_{n}} 0$, implies $T_{n} \xrightarrow{Q_{n}} 0$
(ii) Given $\epsilon>0$, there is a $\delta>0$ s.t. $Q_{n}\left(d P_{n} / d Q_{n}<\delta a_{n}\right)<\epsilon$ f.l.e.n.
(iii) There is $a b>0$ s.t. $\liminf _{n \rightarrow \infty} b a_{n}^{-1} P_{n}\left(d Q_{n} / d P_{n}>b a_{n}^{-1}\right)=1$
(iv) Given $\epsilon>0$, there is a $c>0$ such that $\left\|Q_{n}-Q_{n} \wedge c a_{n}^{-1} P_{n}\right\|<\epsilon$
(v) Under $Q_{n}$, every subsequence of $\left(a_{n}\left(d P_{n} / d Q_{n}\right)^{-1}\right)$ has a further subsequence that converges in $\mathscr{T}_{C}$.

## The model as a uniform space

Take $\mathscr{X}$ a separable metrizable space, with Borel $\sigma$-algebra $\mathscr{B}$.

The class $\mathscr{F}_{n}$ contains all bounded, $\mathscr{B}^{n}$-measurable $f: \mathscr{X}^{n} \rightarrow \mathbb{R}$.

For every $n \geq 1$ and $f \in \mathscr{F}_{n}$, define the entourage,

$$
W_{n, f}=\left\{(P, Q) \in \mathscr{P} \times \mathscr{P}:\left|P^{n} f-Q^{n} f\right|<1\right\} .
$$

Defines uniformity $\mathscr{U}_{n}$ (with topology $\mathscr{T}_{n}$ ). Take $\mathscr{U}_{\infty}=\cup_{n \geq 1} \mathscr{U}_{n}$.

$$
P \rightarrow Q \text { in } \mathscr{T}_{\infty} \quad \Leftrightarrow \quad \int f d P^{n} \rightarrow \int f d Q^{n}
$$

for all $n \geq 1$ and all $f \in \mathscr{F}_{n}$. Note also,

$$
\mathscr{U}_{C} \subset \mathscr{U}_{1} \subset \cdots \subset \mathscr{U}_{\infty} \subset \mathscr{U}_{T V} .
$$

## The Dunford-Pettis theorem

Theorem 25.1 (Dunford-Pettis) Assume $\mathscr{P}$ is dominated by a probability measure $Q$ with densities in $\mathscr{P}_{Q} \subset L^{1}(Q) ; \mathscr{P}_{Q}$ is relatively weakly compact, if and only if, for every $\epsilon>0$ there is an $M>0$ such that,

$$
\sup _{P \in \mathscr{P}} \int_{\{d P / d Q>M\}} \frac{d P}{d Q} d Q<\epsilon
$$

that is, $\mathscr{P}_{Q}$ is uniformly $Q$-integrable.

## Uniform separation

Definition 26.1 Subsets $B, V \subset \mathscr{P}$ are uniformly separated by $\mathscr{U}_{\infty}$, if there exists an entourage $W \in \mathscr{U}_{\infty}$ such that,

$$
(B \times V \cup V \times B) \cap W=\varnothing
$$

In other words, there are $J, m \geq 1, \epsilon>0$ and bounded, measurable functions $f_{1}, \ldots, f_{J}: \mathscr{X}^{m} \rightarrow[0,1]$ such that, for any $P, Q \in B \cup V$, if,

$$
\max _{1 \leq j \leq J}\left|P^{m} f_{j}-Q^{m} f_{j}\right|<\epsilon
$$

then either $P, Q \in B$, or $P, Q \in V$. (If the model is $\mathscr{T}_{\infty}$-compact, $m=1$ suffices).

## The Le Cam-Schwartz theorem

Theorem 27.1 (Le Cam-Schwartz, 1960) Let $\mathscr{P}$ be a model for i.i.d. data $X^{n}$ with disjoint subsets $B, V$. The folllowing are equivalent:
$i$. there exist (uniformly) consistent tests for $B$ vs $V$,
ii. there is a sequence of $\mathscr{U}_{\infty}$-uniformly continuous $\psi_{n}: \mathscr{P} \rightarrow[0,1]$,

$$
\begin{equation*}
\psi_{n}(P) \rightarrow 1_{V}(P) \tag{6}
\end{equation*}
$$

(uniformly) for all $P \in B \cup V$.

## Example: how many clusters?

Observe i.i.d. $X^{n} \sim P^{n}$, where $P$ dominated with density $p$.

Clusters Family $\mathscr{F}$ of kernels $\varphi_{\theta}: \mathbb{R} \rightarrow[0, \infty)$, with parameter $\theta \in \Theta$. Assume $\Theta$ compact and,

$$
\theta \mapsto \int f(x) \varphi_{\theta}(x) d x
$$

is continuous, for every bounded, measurable $f: \mathbb{R} \rightarrow \mathbb{R}$. Define $\Theta_{M}^{\prime}=\Theta^{M} / \sim$.

Model Assume that there is an $M>0$ such that $p$ can be written as,

$$
p_{\lambda, \theta}(x)=\sum_{m=1}^{M} \lambda_{m} p_{\theta_{m}}(x),
$$

for some $M \geq 1$, with $\lambda \in S_{M}=\left\{\lambda \in[0,1]^{M}: \sum_{m} \lambda_{m}=1\right\}, \theta \in \Theta_{M}^{\prime}$.

## Example: how many clusters? (II)

Assume $M$ less than some known $M^{\prime}$. Choose prior $\Pi_{\lambda, M}$ for $\lambda \in S_{M}$ such that, for some $\epsilon>0$,

$$
\Pi_{\lambda, M}\left(\lambda \in S_{M}: \epsilon<\min \left\{\lambda_{m}\right\}, \max \left\{\lambda_{m}\right\}<1-\epsilon\right)=1
$$

For $\theta \in \Theta_{M}^{\prime}$ also choose a prior $\Pi_{\theta, M}$ that 'stays away from the edges'. Define,

$$
\Pi=\sum_{M=1}^{M^{\prime}} \mu_{M} \Pi_{\lambda, M} \times \Pi_{\theta, M}
$$

(for $\sum_{M} \mu_{M}=1$ ).

If $\Pi$ is a KL-prior, posterior odds select the correct number of clusters $M$. If there are no $M^{\prime}$ and $\epsilon$ known, there are sequences $M_{n}^{\prime} \rightarrow \infty$ and $\epsilon_{n} \downarrow 0$ with priors $\Pi_{n}$ that finds the correct number of clusters.

